

Rounding global correlations

This lecture continues our discussion of Sum-of-Squares (SoS) relaxations and how to *round* their solutions into actual assignments for constraint satisfaction problems. The focus is on *dense* Max-2-CSP instances, where the constraint graph has $\Theta(n^2)$ edges. In this regime, obtaining an *additive* approximation guarantee (e.g. within εn^2) is meaningful and often achievable with relatively low levels of the SoS hierarchy.

A low-degree SoS solution gives us *locally consistent* distributions on small sets of variables. A very tempting (but generally incorrect) idea is to sample each variable independently from its one-variable marginal. The loss compared to the SoS objective comes exactly from *global correlations*: if the pairwise marginals μ_{ij} are far from products $\mu_i \times \mu_j$, then independent rounding can destroy value. The main technique in this lecture is to measure correlations by *mutual information* and to iteratively *condition* on carefully chosen variables until the remaining variables have small average mutual information. At that point, naive independent rounding is provably good.

14.1 Dense Max-2-CSP and the SoS relaxation

A *Max-2-CSP* instance is specified by a graph $G = ([n], E)$ and, for every edge $(i, j) \in E$, a constraint function

$$\phi_{ij} : \Sigma \times \Sigma \rightarrow [0, 1].$$

An assignment is $x \in \Sigma^n$, and its value is the total satisfied weight

$$\text{val}(x) = \sum_{(i,j) \in E} \phi_{ij}(x_i, x_j). \quad (14.1)$$

We write $\text{OPT} = \max_{x \in \Sigma^n} \text{val}(x)$. The instance is *dense* if $|E| = cn^2$ for some constant $c > 0$. In this setting, $\text{OPT} = \Theta(n^2)$ and we aim for an additive εn^2 approximation.

THEOREM 14.1 (Dense Max-2-CSP via SoS (informal)). *Let $q = |\Sigma|$. There is a degree $d = O(\log q/\varepsilon^2)$ SoS-based algorithm that outputs an assignment x with*

$$\mathbb{E}[\text{val}(x)] \geq \text{OPT} - \varepsilon|E| = \text{OPT} - O(\varepsilon n^2).$$

To write an SoS relaxation, introduce indicator variables

$$Y_{i,a} \in \{0, 1\} \quad (i \in [n], a \in \Sigma),$$

intended to represent “ $x_i = a$ ”. A valid assignment satisfies the polynomial constraints

$$Y_{i,a}^2 - Y_{i,a} = 0 \quad (\text{Booleanity}) \quad (14.2)$$

$$\sum_{a \in \Sigma} Y_{i,a} = 1 \quad (\text{unique label}). \quad (14.3)$$

In this encoding, the objective (14.1) becomes the degree-2 polynomial

$$P(Y) = \sum_{(i,j) \in E} \sum_{a,b \in \Sigma} \phi_{ij}(a,b) Y_{i,a} Y_{j,b}. \quad (14.4)$$

The degree- d SoS relaxation computes the optimum $\text{SDP}_d(I)$ via the following *feasibility / binary-search* formulation. For a threshold $t \in \mathbb{R}$, consider the feasibility problem: Find a degree- d p.d. μ , s.t.

$$\begin{cases} \tilde{\mathbb{E}}[1] = 1, \\ \tilde{\mathbb{E}}[P(Y)] \geq t, \\ \tilde{\mathbb{E}} \text{ satisfies (14.2)–(14.3) up to degree } d. \end{cases} \quad (14.5)$$

One defines

$$\text{SDP}_d(I) := \sup\{t \in \mathbb{R} : (14.5) \text{ is feasible}\}.$$

The value $\text{SDP}_d(I)$ can be computed (to any precision δ) by binary search on t , each iteration solving the SDP feasibility problem (14.5); this takes time $n^{O(d)} \cdot \text{poly}(\log 1/\delta)$. Any real assignment $x \in \Sigma^n$ gives a feasible pseudo-expectation, which achieves $\tilde{\mathbb{E}}[P(Y)] = \text{val}(x)$; hence $\text{SDP}_d(I) \geq \text{OPT}$.

14.2 Encoding polynomial constraints in SoS

The SoS formalism enforces algebraic constraints by requiring that they hold after multiplying by all low-degree test polynomials. The following is the guiding principle.

FACT 14.2 (How SoS encodes equalities and inequalities). *Let $\tilde{\mathbb{E}}$ be a degree- d pseudo-expectation.*

1. If a true distribution μ is supported on $\{x : P_0(x) = 0\}$, then

$$\mathbb{E}_\mu[P_0(x) Q(x)] = 0 \quad \forall Q.$$

In SoS we relax this to

$$\tilde{\mathbb{E}}[P_0(x) Q(x)] = 0 \quad \forall Q \text{ with } \deg(P_0 Q) \leq d.$$

2. If a true distribution μ is supported on $\{x : P_1(x) \geq 0\}$, then

$$\mathbb{E}_\mu[P_1(x) Q(x)^2] \geq 0 \quad \forall Q.$$

In SoS we relax this to

$$\tilde{\mathbb{E}}[P_1(x) Q(x)^2] \geq 0 \quad \forall Q \text{ with } \deg(P_1 Q^2) \leq d.$$

Proof that these constraints can be enforced by an SDP. We work with a single variable set $x = (x_1, \dots, x_n)$ taking values in a finite alphabet Σ . Index all monomials $x^\alpha = \prod_i x_i^{\alpha_i}$ of degree at most $d/2$ by a set \mathcal{M} . A degree- d pseudo-expectation $\tilde{\mathbb{E}}$ is determined by its *moment vector*

$$\mathbf{m}_\alpha := \tilde{\mathbb{E}}[x^\alpha], \quad |\alpha| \leq d.$$

Moment matrix (base positivity). The fundamental SDP variable is the *moment matrix* $M \in \mathbb{R}^{\mathcal{M} \times \mathcal{M}}$ defined by

$$M_{\alpha,\beta} := \tilde{\mathbb{E}}[x^{\alpha+\beta}], \quad |\alpha|, |\beta| \leq d/2.$$

For any polynomial $Q(x) = \sum_\alpha c_\alpha x^\alpha$ of degree at most $d/2$,

$$Q^\top M Q = \sum_{\alpha,\beta} c_\alpha c_\beta M_{\alpha,\beta} = \tilde{\mathbb{E}} \left[\left(\sum_\alpha c_\alpha x^\alpha \right)^2 \right] = \tilde{\mathbb{E}}[Q(x)^2].$$

Hence the pseudo-expectation axiom $\tilde{\mathbb{E}}[Q^2] \geq 0$ for all Q of degree $\leq d/2$ is *exactly* the constraint $M \succeq 0$ (positive semidefiniteness). This is a standard SDP constraint.

Equality constraints (item 1). Write $P_0(x) = \sum_\gamma p_\gamma x^\gamma$. For any monomial x^β with $\deg(P_0) + |\beta| \leq d$, the constraint $\tilde{\mathbb{E}}[P_0(x) \cdot x^\beta] = 0$ becomes the linear equation

$$\sum_\gamma p_\gamma \mathbf{m}_{\gamma+\beta} = 0.$$

Ranging over all such β yields a finite set of *linear equality constraints* on the moment vector $(\mathbf{m}_\alpha)_{|\alpha| \leq d}$. Since any polynomial Q of the appropriate degree is a linear combination of monomials, $\tilde{\mathbb{E}}[P_0 Q] = 0$ for all such Q is equivalent to this finite set of linear equations, which is an affine constraint in the SDP.

Inequality constraints (item 2). Write $P_1(x) = \sum_\gamma p_\gamma x^\gamma$ and suppose $\deg(P_1) = k$. Define the *localizing matrix* $M^{P_1} \in \mathbb{R}^{\mathcal{M}' \times \mathcal{M}'}$, where \mathcal{M}' indexes monomials of degree at most $(d-k)/2$, by

$$M_{\alpha,\beta}^{P_1} := \tilde{\mathbb{E}}[P_1(x) x^{\alpha+\beta}] = \sum_\gamma p_\gamma \mathbf{m}_{\gamma+\alpha+\beta}.$$

Each entry is a linear function of the moment vector. For any polynomial $Q(x) = \sum_\alpha c_\alpha x^\alpha$ of degree at most $(d-k)/2$,

$$Q^\top M^{P_1} Q = \sum_{\alpha,\beta} c_\alpha c_\beta M_{\alpha,\beta}^{P_1} = \tilde{\mathbb{E}}[P_1(x) \cdot Q(x)^2].$$

Hence the condition $\tilde{\mathbb{E}}[P_1(x) Q(x)^2] \geq 0$ for all such Q is exactly $M^{P_1} \succeq 0$, another positive-semidefiniteness constraint in the SDP.

Summary. The degree- d SoS feasibility problem (14.5) can be written explicitly as:

$$\text{find } \mathbf{m} = (\mathbf{m}_\alpha)_{|\alpha| \leq d} \in \mathbb{R}^{\mathcal{M} \cup \{0\}} \text{ such that } M(\mathbf{m}) \succeq 0, \quad M^{P_1}(\mathbf{m}) \succeq 0, \quad L(\mathbf{m}) = 0, \quad \mathbf{m}_0 = 1,$$

where M is the moment matrix, M^{P_1} is a localizing matrix for each inequality constraint, and $L(\mathbf{m}) = 0$ encodes the equality constraints and the objective threshold $\tilde{\mathbb{E}}[P(Y)] \geq t$. The number of variables is $|\mathcal{M}| = n^{O(d)}$, and each matrix has size at most $n^{O(d)} \times n^{O(d)}$, so the SDP can be solved in time $n^{O(d)}$. \square

In our Max-2-CSP relaxation, the constraints (14.2) and (14.3) are equalities, so we require the corresponding identities to hold in the sense of item (1) above.

14.3 Reweighting and pseudo-conditioning

A powerful feature of pseudo-distributions is that we can mimic basic probabilistic operations, such as reweighting and conditioning, at low degree.

DEFINITION 14.3 (Reweighting). Let $\tilde{\mathbb{E}}$ be a degree- d pseudo-expectation and let p be a degree- k *sum-of-squares* polynomial (so $p = \sum_r r(x)^2$) with $\tilde{\mathbb{E}}[p] > 0$. Define the reweighted pseudo-expectation $\tilde{\mathbb{E}}_p$ on polynomials Q of degree at most $d - k$ by

$$\tilde{\mathbb{E}}_p[Q] = \frac{\tilde{\mathbb{E}}[p \cdot Q]}{\tilde{\mathbb{E}}[p]}. \quad (14.6)$$

FACT 14.4. *If $\tilde{\mathbb{E}}$ is degree- d and p has degree k , then $\tilde{\mathbb{E}}_p$ is a degree- $(d-k)$ pseudo-expectation.*

Proof. **Normalization.**

$$\tilde{\mathbb{E}}_p[1] = \frac{\tilde{\mathbb{E}}[p \cdot 1]}{\tilde{\mathbb{E}}[p]} = \frac{\tilde{\mathbb{E}}[p]}{\tilde{\mathbb{E}}[p]} = 1.$$

Positivity. Let Q be a polynomial with $\deg(Q) \leq (d - k)/2$, so that in particular $\deg(Q^2) \leq d - k$. Write $p = \sum_r r^2$ where each r has degree at most $k/2$. Then rQ has degree $\deg(r) + \deg(Q) \leq k/2 + (d - k)/2 = d/2$. Hence each $(rQ)^2$ has degree at most d , and $\tilde{\mathbb{E}}[(rQ)^2] \geq 0$ by the positivity of $\tilde{\mathbb{E}}$. Now expand:

$$\tilde{\mathbb{E}}_p[Q^2] = \frac{\tilde{\mathbb{E}}[p \cdot Q^2]}{\tilde{\mathbb{E}}[p]} = \frac{\tilde{\mathbb{E}}[(\sum_r r^2)Q^2]}{\tilde{\mathbb{E}}[p]} = \frac{\sum_r \tilde{\mathbb{E}}[r^2 Q^2]}{\tilde{\mathbb{E}}[p]} = \frac{\sum_r \tilde{\mathbb{E}}[(rQ)^2]}{\tilde{\mathbb{E}}[p]} \geq 0,$$

where the last step uses $\tilde{\mathbb{E}}[(rQ)^2] \geq 0$ for each r and $\tilde{\mathbb{E}}[p] > 0$ by assumption.

Constraint satisfaction. If $\tilde{\mathbb{E}}$ satisfies the equality constraint $\tilde{\mathbb{E}}[P_0 Q'] = 0$ for all Q' with $\deg(P_0 Q') \leq d$, then for any polynomial Q'' with $\deg(P_0 Q'') \leq d - k$,

$$\tilde{\mathbb{E}}_p[P_0 Q''] = \frac{\tilde{\mathbb{E}}[p \cdot P_0 Q'']}{\tilde{\mathbb{E}}[p]} = 0,$$

since $\deg(p \cdot P_0 Q'') \leq k + (d - k) = d$. Thus $\tilde{\mathbb{E}}_p$ inherits all constraints of $\tilde{\mathbb{E}}$ up to the reduced degree $d - k$. \square

Reweighting recovers ordinary conditioning in the Boolean setting. For a true distribution on $x_n \in \{0, 1\}$,

$$\mu \mid x_n = 1 \propto \mathbf{1}(x_n = 1) \mu.$$

On $\{0, 1\}$ we have $\mathbf{1}(x_n = 1) = x_n = x_n^2$, so a pseudo-version of conditioning on $x_n = 1$ is obtained by reweighting with $p(x) = x_n^2$:

$$\tilde{\mathbb{E}}_{\mu \mid x_n=1}[Q] = \frac{\tilde{\mathbb{E}}_{\mu}[Q \cdot x_n^2]}{\tilde{\mathbb{E}}_{\mu}[x_n^2]}. \quad (14.7)$$

Similarly, conditioning on $x_n = 0$ corresponds to reweighting with $(1 - x_n)^2$:

$$\tilde{\mathbb{E}}_{\mu \mid x_n=0}[Q] = \frac{\tilde{\mathbb{E}}_{\mu}[Q \cdot (1 - x_n)^2]}{\tilde{\mathbb{E}}_{\mu}[(1 - x_n)^2]}. \quad (14.8)$$

In our Max-2-CSP encoding, conditioning on “ $x_i = a$ ” is implemented by reweighting with $Y_{i,a}^2$ (which equals $Y_{i,a}$ on true assignments). Each such conditioning consumes degree 2.

14.4 Local distributions from low-degree moments

A key intuition is that a degree- d pseudo-distribution behaves like a real distribution on any set of at most $d/2$ variables.

FACT 14.5 (Local distributions). *Let $\tilde{\mathbb{E}}$ be a degree- d pseudo-expectation satisfying (14.2)–(14.3). Fix indices i_1, \dots, i_k with $k \leq d/2$ and labels $a_1, \dots, a_k \in \Sigma$. Define*

$$\tilde{\Pr}[X_{i_1} = a_1, \dots, X_{i_k} = a_k] := \tilde{\mathbb{E}} \left[\prod_{t=1}^k Y_{i_t, a_t} \right].$$

Then these numbers form a valid probability distribution on Σ^k (nonnegative and summing to 1).

*Proof. **Nonnegativity.*** Since $k \leq d/2$, the product $\prod_{t=1}^k Y_{i_t, a_t}$ has degree $k \leq d/2$. By the Booleanity constraints (14.2), $Y_{i,a}^2 = Y_{i,a}$ holds as a polynomial identity modulo the ideal, so inside the pseudo-expectation:

$$\tilde{\mathbb{E}} \left[\prod_{t=1}^k Y_{i_t, a_t} \right] = \tilde{\mathbb{E}} \left[\prod_{t=1}^k Y_{i_t, a_t}^2 \right].$$

The right-hand side has degree $2k \leq d$ and factors as a square $\tilde{\mathbb{E}} \left[\left(\prod_{t=1}^k Y_{i_t, a_t} \right)^2 \right] \geq 0$, where positivity follows from the pseudo-expectation axiom (the squared polynomial has degree $2k \leq d$).

Normalization. Sum over all label tuples. The linearity of $\tilde{\mathbb{E}}$ and the unique-label constraints (14.3) give

$$\begin{aligned} \sum_{a_1 \in \Sigma} \cdots \sum_{a_k \in \Sigma} \tilde{\mathbb{E}} \left[\prod_{t=1}^k Y_{i_t, a_t} \right] &= \tilde{\mathbb{E}} \left[\sum_{a_1 \in \Sigma} \cdots \sum_{a_k \in \Sigma} \prod_{t=1}^k Y_{i_t, a_t} \right] \\ &= \tilde{\mathbb{E}} \left[\prod_{t=1}^k \left(\sum_{a_t \in \Sigma} Y_{i_t, a_t} \right) \right] \\ &= \tilde{\mathbb{E}} \left[\prod_{t=1}^k 1 \right] = \tilde{\mathbb{E}}[1] = 1, \end{aligned}$$

where the interchange of sums and the product is valid because the i_t are distinct, so the sums over different indices act on disjoint sets of variables. \square

In particular, if $d \geq 4$ then we obtain genuine one-variable marginals μ_i and pairwise marginals μ_{ij} (distributions on Σ and Σ^2 , respectively).

14.5 Naive rounding and mutual information

Let $\tilde{\mathbb{E}} = \tilde{\mathbb{E}}_\mu$ be the degree- d SoS solution returned by the SDP. Assume $d \geq 4$ so that the one- and two-variable marginals are well defined. Let X_i denote the random label of variable i under these marginals:

$$\mu_i(a) = \tilde{\mathbb{E}}[Y_{i,a}], \quad \mu_{ij}(a,b) = \tilde{\mathbb{E}}[Y_{i,a}Y_{j,b}].$$

Define the SoS objective value

$$\begin{aligned} \text{SDP} &:= \tilde{\mathbb{E}} \left[\sum_{(i,j) \in E} \sum_{a,b \in \Sigma} \phi_{ij}(a,b) Y_{i,a} Y_{j,b} \right] \\ &= \sum_{(i,j) \in E} \mathbb{E}_{(X_i, X_j) \sim \mu_{ij}} [\phi_{ij}(X_i, X_j)]. \end{aligned} \quad (14.9)$$

A *naive rounding* samples each X_i *independently* from μ_i and outputs the resulting assignment. Its expected value is

$$\text{naive} = \sum_{(i,j) \in E} \mathbb{E}_{X_i \sim \mu_i, X_j \sim \mu_j} [\phi_{ij}(X_i, X_j)]. \quad (14.10)$$

The gap between (14.9) and (14.10) is controlled by how far μ_{ij} is from the product $\mu_i \times \mu_j$.

CLAIM 14.6 (Correlation loss bounded by total variation). *For each edge $(i, j) \in E$,*

$$\mathbb{E}_{(X_i, X_j) \sim \mu_{ij}} [\phi_{ij}(X_i, X_j)] - \mathbb{E}_{X_i \sim \mu_i, X_j \sim \mu_j} [\phi_{ij}(X_i, X_j)] \leq \|\mu_{ij} - \mu_i \times \mu_j\|_{\text{TV}}.$$

Consequently,

$$\text{naive} \geq \text{SDP} - \sum_{(i,j) \in E} \|\mu_{ij} - \mu_i \times \mu_j\|_{\text{TV}}. \quad (14.11)$$

Proof. Per-edge bound. Recall that for any two distributions P, Q on a finite set \mathcal{X} and any function $f : \mathcal{X} \rightarrow [0, 1]$,

$$|\mathbb{E}_P[f] - \mathbb{E}_Q[f]| = \left| \sum_x f(x)(P(x) - Q(x)) \right| \leq \sum_x |P(x) - Q(x)| = 2 \|P - Q\|_{\text{TV}}.$$

Since $\phi_{ij} \in [0, 1]$ and we are bounding from below:

$$\mathbb{E}_{(X_i, X_j) \sim \mu_{ij}} [\phi_{ij}(X_i, X_j)] - \mathbb{E}_{X_i \sim \mu_i, X_j \sim \mu_j} [\phi_{ij}(X_i, X_j)] \leq |\mathbb{E}_{\mu_{ij}}[\phi_{ij}] - \mathbb{E}_{\mu_i \times \mu_j}[\phi_{ij}]| \leq \|\mu_{ij} - \mu_i \times \mu_j\|_{\text{TV}}.$$

(The factor of 2 disappears because $f \in [0, 1]$ gives a sharper bound of $\|P - Q\|_{\text{TV}}$ by the dual characterization $\|P - Q\|_{\text{TV}} = \sup_{f: 0 \leq f \leq 1} (\mathbb{E}_P[f] - \mathbb{E}_Q[f])$.)

Summing over edges. Substituting into the definitions (14.9) and (14.10):

$$\begin{aligned} \text{naive} &= \sum_{(i,j) \in E} \mathbb{E}_{X_i \sim \mu_i, X_j \sim \mu_j} [\phi_{ij}(X_i, X_j)] \\ &\geq \sum_{(i,j) \in E} (\mathbb{E}_{\mu_{ij}}[\phi_{ij}(X_i, X_j)] - \|\mu_{ij} - \mu_i \times \mu_j\|_{\text{TV}}) \\ &= \text{SDP} - \sum_{(i,j) \in E} \|\mu_{ij} - \mu_i \times \mu_j\|_{\text{TV}}. \quad \square \end{aligned}$$

To turn (14.11) into an additive εn^2 bound, we need the *average* pairwise total variation distance to be $O(\varepsilon)$. This is where mutual information enters.

FACT 14.7 (Pinsker + mutual information). *For distributions P, Q on the same space,*

$$\|P - Q\|_{\text{TV}}^2 \leq \frac{1}{2} \text{KL}(P\|Q).$$

In particular, for a pair $(X_i, X_j) \sim \mu_{ij}$,

$$\text{KL}(\mu_{ij} \parallel \mu_i \times \mu_j) = \text{I}(X_i; X_j).$$

Combining Claim 14.6 with Cauchy–Schwarz and Fact 14.7 gives the key bound on naive rounding. Starting from (14.11):

$$\begin{aligned} \text{SDP} - \text{naive} &\leq \sum_{(i,j) \in E} \|\mu_{ij} - \mu_i \times \mu_j\|_{\text{TV}} \\ &\leq |E| \cdot \left(\frac{1}{|E|} \sum_{(i,j) \in E} \|\mu_{ij} - \mu_i \times \mu_j\|_{\text{TV}} \right) \\ &\leq |E| \cdot \sqrt{\mathbb{E}_{(i,j) \in E} \left[\|\mu_{ij} - \mu_i \times \mu_j\|_{\text{TV}}^2 \right]} \end{aligned} \quad (14.12)$$

where the last step is Cauchy–Schwarz (Jensen applied to the concave square root, or equivalently $\frac{1}{m} \sum_k z_k \leq \sqrt{\frac{1}{m} \sum_k z_k^2}$). By Pinsker’s inequality (Fact 14.7), $\|\mu_{ij} - \mu_i \times \mu_j\|_{\text{TV}}^2 \leq \frac{1}{2} \text{KL}(\mu_{ij} \parallel \mu_i \times \mu_j) = \frac{1}{2} \text{I}(X_i; X_j)$. Substituting into (14.12):

$$\begin{aligned} \text{naive} &\geq \text{SDP} - |E| \cdot \sqrt{\mathbb{E}_{(i,j) \in E} \left[\|\mu_{ij} - \mu_i \times \mu_j\|_{\text{TV}}^2 \right]} \\ &\geq \text{SDP} - |E| \cdot \sqrt{\frac{1}{2} \mathbb{E}_{(i,j) \in E} [\text{I}(X_i; X_j)]}. \end{aligned} \quad (14.13)$$

Thus, if the *average mutual information*

$$A := \mathbb{E}_{(i,j) \in E} [\text{I}(X_i; X_j)]$$

satisfies $A \leq 2\varepsilon^2$, then

$$\text{naive} \geq \text{SDP} - \varepsilon|E| \geq \text{OPT} - \varepsilon|E|.$$

When A is not small, we will *condition* to reduce it.

14.6 Correlation rounding via entropy potential

We now describe the “rounding by global correlations” procedure.

Let $\mu^{(0)} = \mu$ be the degree- d pseudo-distribution returned by SoS, and set $t = 0$. At each stage we maintain: (i) a set of already fixed variables, and (ii) a conditioned pseudo-distribution $\mu^{(t)}$ (implemented via reweighting as in Section 14.3) over the remaining variables.

Rounding algorithm. Let A_t be the average mutual information among *alive* (unfixed) variables under $\mu^{(t)}$.

1. If $A_t \leq \varepsilon^2$, perform naive independent rounding on all remaining variables using their one-variable marginals under $\mu^{(t)}$ and output the full assignment.
2. Otherwise, choose an alive index i such that

$$\mathbb{E}_{j \text{ alive}, j \neq i} [\mathbb{I}_{\mu^{(t)}}(X_i; X_j)] \geq \varepsilon^2.$$

(Existence follows by averaging.) Sample a label $a \sim \mu_i^{(t)}$ and fix $X_i = a$. Update $\mu^{(t+1)}$ by pseudo-conditioning on the event $Y_{i,a} = 1$.

3. Stop if $t = C \log q / \varepsilon^2$ for a sufficiently large absolute constant C ; otherwise increment t and repeat.

Step (1) is justified by (14.13) applied to $\mu^{(t)}$. The point of steps (2)–(3) is to argue that we reach (1) quickly, before running out of degree.

CLAIM 14.8 (Termination with high probability). *With probability at least $1 - 1/C$, the rounding algorithm reaches step (1) before $t = C \log q / \varepsilon^2$.*

Proof. **Entropy potential.** Consider the potential function

$$\Phi_t := \mathbb{E}_{j \text{ alive}} \mathbb{H}_{\mu^{(t)}}(X_j),$$

the average entropy of the alive variables. Since each X_j takes values in Σ , we have $0 \leq \mathbb{H}(X_j) \leq \log q$, so

$$0 \leq \Phi_t \leq \log q \quad \text{and in particular} \quad \Phi_0 \leq \log q.$$

Expected drop in potential. Suppose the algorithm executes step (2) at time t and picks index i satisfying

$$\mathbb{E}_{j \text{ alive}, j \neq i} \mathbb{I}_{\mu^{(t)}}(X_i; X_j) \geq \varepsilon^2.$$

Such an index exists because if every alive i had $\mathbb{E}_{j \neq i} \mathbb{I}(X_i; X_j) < \varepsilon^2$, then averaging over alive i would give $A_t < \varepsilon^2$ (up to the asymmetry from dropping $j = i$, which contributes 0), contradicting the assumption $A_t > \varepsilon^2$.

After sampling $a \sim \mu_i^{(t)}$ and updating $\mu^{(t+1)} = \mu^{(t)} \mid X_i = a$, note that variable i is fixed and leaves the alive set. The expected drop in potential is

$$\begin{aligned} \mathbb{E}_a[\Phi_t - \Phi_{t+1}] &= \mathbb{E}_{j \text{ alive}, j \neq i} \left(\mathbb{H}_{\mu^{(t)}}(X_j) - \mathbb{E}_{a \sim \mu_i^{(t)}} \mathbb{H}_{\mu^{(t+1)}}(X_j) \right) \\ &= \mathbb{E}_{j \text{ alive}, j \neq i} \left(\mathbb{H}_{\mu^{(t)}}(X_j) - \mathbb{H}_{\mu^{(t)}}(X_j \mid X_i) \right) \\ &= \mathbb{E}_{j \text{ alive}, j \neq i} \mathbb{I}_{\mu^{(t)}}(X_i; X_j) \geq \varepsilon^2, \end{aligned}$$

using the identity $\mathbb{H}(X_j) - \mathbb{H}(X_j \mid X_i) = \mathbb{I}(X_i; X_j)$ and the law of total expectation ($\mathbb{E}_a \mathbb{H}(X_j \mid X_i = a) = \mathbb{H}(X_j \mid X_i)$).

Markov bound. Let T be the (random) number of times step (2) is executed before reaching step (1). By linearity of expectation and the per-step drop:

$$\varepsilon^2 \mathbb{E}[T] \leq \mathbb{E}[\Phi_0 - \Phi_T] \leq \Phi_0 \leq \log q,$$

where we used $\Phi_T \geq 0$. Hence $\mathbb{E}[T] \leq \log q / \varepsilon^2$. Applying Markov's inequality:

$$\mathbb{P} \left[T > C \frac{\log q}{\varepsilon^2} \right] \leq \frac{\mathbb{E}[T]}{C \log q / \varepsilon^2} \leq \frac{1}{C}. \quad \square$$

Degree budget. Each conditioning step (step (2)) uses a degree-2 reweighting polynomial ($Y_{i,a}^2$), so it reduces the available SoS degree by exactly 2. After T conditioning steps the remaining pseudo-distribution has degree $d - 2T$. In the worst case $T = C \log q/\varepsilon^2$ (conditioned on not exceeding the budget), so we need $d - 2C \log q/\varepsilon^2 \geq 4$, i.e. $d \geq 4 + 2C \log q/\varepsilon^2$. Taking $d = \Theta(\log q/\varepsilon^2)$ satisfies this, and the SDP with this degree can be solved in time $n^{O(d)} = n^{O(\log q/\varepsilon^2)}$, which is polynomial in n for fixed q and ε .

Putting it together. At the time step (1) is reached, the remaining pseudo-distribution $\mu^{(t)}$ has degree at least 4, so pairwise marginals are well defined. By construction, the average mutual information $A_t \leq \varepsilon^2$. Applying (14.13) to $\mu^{(t)}$ and noting that the SoS objective can only improve after conditioning (the conditioned pseudo-expectation still satisfies all constraints and gives objective value at least that of the original, since we sample conditionings from the marginals), we obtain

$$\mathbb{E}[\text{val}(\text{output})] \geq \text{SDP}_d(I) - \varepsilon|E| \geq \text{OPT} - \varepsilon|E|.$$

This completes the proof of Theorem 14.1.